# DOMINATION INEQUALITY FOR MARTINGALE TRANSFORMS OF A RADEMACHER SEQUENCE

#### **BY**

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#### ABSTRACT

Let  $f_n = \sum_{k=1}^n v_k r_k$ ,  $n = 1,...$ , be a martingale transform of a Rademacher sequence  $(r_n)$  and let  $(r'_n)$  be an independent copy of  $(r_n)$ . The main result of this paper states that there exists an absolute constant K such that for all  $p, 1 \leq p < \infty$ , the following inequality is true:

$$
\Big\| \sum v_k r_k \Big\|_p \leq K \Big\| \sum v_k r'_k \Big\|_p.
$$

In order to prove this result, we obtain some inequalities which may be of independent interest. In particular, we show that for every sequence of scalars  $(a_n)$  one has

$$
\bigg\|\sum a_kr_k\bigg\|_p\approx K_{1,2}((a_n),\sqrt{p}),
$$

where  $K_{1,2}((a_n), \sqrt{p}) = K((a_n), \sqrt{p}; \ell_1, \ell_2)$  is the K-interpolation norm between  $l_1$  and  $l_2$ . We also derive a new exponential inequality for martingale transforms of a Rademacher sequence.

### 1. Introduction

This paper concerns the *Lp-norm* inequalities for certain martingale transforms. Let  $(\Omega, \mathcal{F}, P)$  be a probability space. Let  $(\mathcal{F}_n)$  be an increasing sequence of sub- $\sigma$ algebras of F. A sequence  $(X_n)$  of random variables is said to be  $(\mathcal{F}_n)$ -predictable (or just predictable, if there is no risk of confusion) provided for each  $n \geq 1$ ,  $X_n$  is

Received August 24, 1992 and in revised form January 12, 1993

<sup>\*</sup> This research was supported in part by an NSF grant and an FRPD grant at NCSU.

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 $\mathcal{F}_{n-1}$ -measurable random variable. All equalities or inequalities between random variables are assumed to hold almost surely. For any sequence  $(X_n)$  of random variables we shall write  $X^* = \sup_n |X_n|$  and  $X_n^* = \max_{1 \leq k \leq n} |X_k|$ . The  $L_p$ -norm,  $1 \leq p \leq \infty$ , of a random variable X is denoted by  $||X||_p$ , and  $I(A)$  or  $I_A$  will denote the indicator function of a set A. Let  $({\xi_n})$  be a sequence of independent, mean-zero random variables and set  $\mathcal{F}_n = \sigma(\xi_1,\ldots,\xi_n), n = 1, \ldots$  If  $(v_n)$  is a sequence predictable relative to  $(\mathcal{F}_n)$  then a sequence  $(f_n)$  defined by

$$
f_n=\sum_{k=1}^n v_k \xi_k, \qquad n=1,\ldots
$$

is a martingale and is referred to as martingale transform of the sequence  $(\xi_k)$ . Denote by  $({\xi'_n})$  an independent copy of a sequence  $({\xi_n})$ . It is known that for each  $p, 1 \leq p < \infty$ , there exist constants  $C_p$  and  $K_p$  depending only on p such that for all predictable sequences  $(v_n)$  of random variables the following two-sided estimate holds:

(\*) 
$$
C_p^{-1} \|\sum_{k=1}^n v_k \zeta'_k\|_p \leq \|\sum_{k=1}^n v_k \zeta_k\|_p \leq K_p \|\sum_{k=1}^n v_k \zeta'_k\|_p, \quad n = 1, ...
$$

Several proofs of (\*) as well as of more general results axe available (cf. e.g. Zinn [26], Hitczenko [10, 13] or McConnell [23]). The proofs given in the last two papers show that both constants  $C_p$  and  $K_p$  are bounded above by  $O(p)$  for large p. It is not hard to see that the above bound on  $C_p$  is optimal i.e. that there are sequences  $({\xi_n})$  and  $(v_n)$  for which:

$$
\Big\| \sum_{k=1}^n v_k \zeta'_k \Big\|_p \approx O(p) \Big\| \sum_{k=1}^n v_k \zeta_k \Big\|_p.
$$

Indeed, if  $(r_n)$  is a Rademacher sequence (i.e. a sequence of i.i.d. random variables such that  $P(r_n = 1) = P(r_n = -1) = 1/2$  and  $v_k = I(\tau \ge k)$ , where  $\tau = \inf\{n \in \mathbb{N}: r_n = r_{n-1}\},\$  then  $\|\sum v_k r_k\|_p \leq 1$ . On the other hand, denoting by  $[x]$  the integer part of a number x, we have that

$$
\|\sum v_k r'_k\|_p^p \ge \sum_{k=1}^{\infty} P(\tau = k) E \Big| \sum_{j=1}^k r'_j \Big|^p \ge P(\tau = [p]) E \Big| \sum_{j=1}^{[p]} r'_j \Big|^p
$$
  
 
$$
\ge P(\tau = [p])[p]^p P(r'_j = 1, j = 1, ..., [p]) \approx (p/4)^p.
$$

In this paper we focus our attention on the right inequality in  $(*)$ , which is more interesting from the point of view of applications. Our main result asserts that, if  $\xi_k = r_k$  then the constant  $K_p$  may actually be chosen to be *independent* of p. This should be compared with a result of Klass [16, Theorem 3.1] who showed that if  $({\xi_k})$  is a sequence of independent random variables and  $v_k = I(\tau \geq k)$ , where  $\tau$  is a stopping time with respect to  $(\xi_k)$ , then  $(*)$  holds with constant  $K_p$ independent of p.

Our proof relies on martingale methods, and we believe that some aspects of our approach may be of independent interest. In particular, we obtain a two-sided inequality for the  $L_p$ -norm of the sum  $\sum a_k r_k$  in terms of  $K((a_n), \sqrt{p}; \ell_1, \ell_2)$ ; the K-interpolation norm of a sequence  $(a_n)$ . We also establish a new exponential inequality for martingale transforms of a Rademacher sequence.

In view of the main result of this paper one may be tempted to ask whether the following statement is true: There exists a constant  $C$  such that for all sequences  $(v_n)$  predictable with respect to the natural filtration generated by a Rademacher sequence  $(r_n)$  and all  $t > 0$  one has

$$
P(|\sum v_k r_k| \ge Ct) \le CP(|\sum v_k r'_k| \ge t),
$$

where  $(r'_n)$  is an independent copy of  $(r_n)$  (such inequality was actually conjectured by Kwapień and Woyczyński [19]). A negative answer to this question was given recently by Talagrand (personal communication). We shall present his example and would like to thank him for permitting us to include it in this paper.

The paper is organized as follows. The next section contains some preliminary lemmas. Sections 3 and 4 contain the main ingredients needed for the proof. In the former we prove an approximation for the  $L_p$ -norm of Rademacher averages, while in the latter we obtain an exponential inequality for martingale transforms of  $(r_n)$ . The main result is stated in Section 5. It may be easily generalized to martingale transforms satisfying certain regularity condition. This extension is not completely satisfactory, since the constant  $K$  depends on the regularity constant of the original martingale. Lastly, in Section 6, we present Talagrand's example.

## 2. Preliminaries

As we mentioned in the introduction, we shall use martingale methods. We begin by recalling briefly necessary terminology and we refer the reader to the paper Burkholder and Gundy [6] for more details.

Given a martingale  $f = (f_n)$  with difference sequence  $d = (d_n)$  and two stopping times  $\nu$  and  $\mu$  such that  $\nu \geq \mu$ , for  $n = 1, \ldots$ , we let

$$
\mu f_n^{\nu} = \sum_{k=1}^n I(\mu < k \leq \nu) d_k.
$$

Then the sequence  $^{\mu}f^{\nu} = {^{\mu}f^{\nu}_n}$  is also a martingale (referred to as f started at  $\mu$  and stopped at  $\nu$ ). In particular,  $f^n = (f_0, \ldots, f_{n-1}, f_n, f_n, \ldots)$ . Let  $(\mathcal{F}_n)$ be an increasing sequence of  $\sigma$ -algebras and assume that  $\mathcal N$  is a collection of martingales relative to  $(\mathcal{F}_n)$  which is closed under starting and stopping (i.e.,  $f \in \mathcal{N}$  implies  $^{\mu}f^{\nu} \in \mathcal{N}$  for all stopping times  $\nu$  and  $\mu$  satisfying  $\nu \geq \mu$ ). We shall consider an operator T defined on a collection  $\mathcal N$  with values in the class of all nonnegative random variables on the same probability space. We shall assume that  $T$  satisfies the following conditions (cf. Burkholder and Gundy  $[6]$ ):

- (B1) T is quasi-linear, i.e.,  $T(f + g) \leq \gamma(T(f) + T(g))$ , for some nonnegative  $\gamma$  and all martingales  $f, g \in \mathcal{N}'$ .
- (B2) T is local, i.e.,  $T(f) = 0$  on the set  $\{s(f) = 0\}, f \in \mathcal{N}$  where

$$
s(f) = \left(\sum E(d_k^2|\mathcal{F}_{k-1})\right)^{1/2}
$$

is the conditional square function of f.

(B3) T is symmetric, i.e.,  $T(f) = T(-f)$ , for all martingales  $f \in \mathcal{N}$ . If  $\gamma = 1$  then T is sublinear. An operator T is called measurable (resp. predictable) if, for  $n = 1, ..., T(f<sup>n</sup>)$  is  $\mathcal{F}_n$ -measurable (resp.  $\mathcal{F}_{n-1}$ -measurable) random variable. For example, the square function  $S(f)$  defined by  $S(f) = (\sum d_k^2)^{1/2}$ is a measurable operator while the conditional square function is a predictable operator on the collection of all martingales relative to  $(\mathcal{F}_n)$ .

Our first lemma gives a sufficient condition for a comparison of the  $L_p$ norms of random variables and can be found in Burkholder [4, Lemma 7.1]:

LEMMA 2.1: Fix  $0 < p < \infty$ . Suppose that X and Y are nonnegative random *variables and that there exist positive numbers*  $\delta$ *,*  $\beta > 1 + \delta$  *and*  $\epsilon < 1/\beta^p$  *such that for all positive*  $\lambda$ 's one has

$$
P(X \geq \beta \lambda, Y < \delta \lambda) \leq \epsilon P(X \geq \lambda).
$$

*Then* 

$$
||X||_p^p \leq \frac{(\beta/\delta)^p}{1-\beta^p\epsilon}||Y||_p^p.
$$

The above lemma in a more general form was originally used to prove some martingale inequalities (cf. Burkholder [4], Burkholder and Gundy [6] and Burkholder, Davis and Gundy [5]. It was realized later (see e.g. Bafiuelos [1] or Hitczenko [11, 12]) that it could also be used to give rather precise information on the growth rate of the constants involved in some of those inequalities. The next lemma hints at the way we are going to use that idea.

LEMMA 2.2: Fix  $1 \leq p < \infty$ . Let N be any class of martingales which is closed under starting and *stopping and let T be a predictable operator on*  $N$  satisfying *conditions* (B1)–(B3). *Suppose that there exist positive constants*  $c_0$  and  $\kappa$  such that, for all mean-zero martingales  $f \in \mathcal{N}$  the following inequality holds:

$$
P(|f_n| > c||T^*(f^n)||_{\infty}) \leq 2 \exp \left\{-\kappa c^2 p\right\}, \quad c \geq c_0
$$

*(recall that, according to our notation,*  $T^*(f^n) = \max_{1 \leq k \leq n} T(f^k)$ ). Then, *there exists an absolute constant*  $\vartheta = \vartheta(\kappa, \gamma)$  *such that, for all 6 > 0 and*  $\beta$  *>*  $1 + \delta(1 + c_0)$ , we have that

$$
P(f^* \geq \beta \lambda, T^*(f) \vee d^* < \delta \lambda) \leq 2 \exp \left\{-\frac{\vartheta(\beta-1-\delta)^2 p}{\delta^2}\right\} P(f^* \geq \lambda).
$$

*Consequently, if additionally,*  $||d^*||_p \leq \alpha ||T^*(f)||_p$ , for some absolute constant  $\alpha$ *then* 

$$
||f^*||_p \leq C||T^*(f)||_p,
$$

*where C is an absolute constant.* 

*Proof."* Our proof follows usual argument; we let

$$
\nu = \inf\{n \in \mathbb{N} : |f_n| \ge \beta \lambda\},\
$$

$$
\mu = \inf\{n \in \mathbb{N} : |f_n| \ge \lambda\},\
$$

$$
\tau = \inf\{n \in \mathbb{N} : T(f^{n+1}) \vee d_n^* \ge \delta \lambda\}.
$$

Then, on the set  $\{ \nu < \infty, \tau = \infty \}$  we have

$$
\big|\sum_{k=\mu+1}^{\nu\wedge\tau}d_k\big|\geq\big|\sum_{k=1}^{\nu\wedge\tau}d_k\big|-\big|\sum_{k=1}^{\mu-1}d_k\big|-\big|d_\mu\big|\geq(\beta-1-\delta)\lambda.
$$

Moreover, by computation similar to that in Hitczenko [12, proof of Theorem 3.1] or Burkholder and Gundy [6, bottom half of page 256] we infer that  $||T^*(Hf^{\nu\wedge\tau})||_{\infty} \leq 2\gamma^2\delta\lambda$ . Therefore, as long as  $(\beta - 1 - \delta)/2\gamma^2\delta \geq c_0$ , we have

$$
P(f^* \ge \beta \lambda, T^*(f) \vee d^* < \delta \lambda, )
$$
\n
$$
= P(\nu < \infty, \tau = \infty)
$$
\n
$$
\le P(|\sum_{k=\mu}^{\nu \wedge \tau} d_k| > (\beta - 1 - \delta) \lambda)
$$
\n
$$
\le P(|\sum_{k=\mu}^{\nu \wedge \tau} d_k| > \frac{\beta - 1 - \delta}{2\gamma^2 \delta} ||T^*(\mu f^{\nu \wedge \tau})||_{\infty})
$$
\n
$$
= EP(|\sum_{k=\mu}^{\nu \wedge \tau} d_k| > \frac{\beta - 1 - \delta}{2\gamma^2 \delta} ||T^*(\mu f^{\nu \wedge \tau})||_{\infty} |\mathcal{F}_{\mu}).
$$

Note that, conditionally on  $\mathcal{F}_{\mu}$ , " $f^{\nu \wedge \tau}$  is a mean-zero martingale. Therefore, it follows from our assumption (applied to the conditional measure  $P(\cdot | \mathcal{F}_{\mu})$  ) that the conditional probability above does not exceed  $2 \exp \left\{-\frac{\kappa(\beta-1-\delta)^2 p}{(2\gamma^2\delta)^2}\right\}$ , if  $\mu < \infty$ , and is zero if  $\mu = \infty$ . Consequently, the last expectation does not exceed

$$
2E\exp\left\{-\frac{\kappa(\beta-1-\delta)^2p}{\delta^2}\right\}I(\mu<\infty)=2\exp\left\{-\frac{\kappa(\beta-1-\delta)^2p}{\delta^2}\right\}P(f^*>\lambda).
$$

This completes the proof of the first statement. The second statement is now a consequence of Lemma 2.1.  $\blacksquare$ 

We shall now define an operator on martingale transforms of Rademacher sequence. Assume for convenience that  $(r_n)$  and  $(r'_n)$  are defined on different probability spaces,  $(\Omega, \mathcal{F}, P)$  and  $(\Omega', \mathcal{F}', P')$ , respectively. The symbol  $E'(-)$ will be used to denote the integration with respect to the measure  $P'$ . Let  $\mathcal N$ be a class of all martingale transforms of a sequence  $(r_n)$ , i.e.  $\mathcal N$  consists of all martingales  $f = (f_n)$  of the form:

$$
f_n=\sum_{k=1}^n v_k r_k, \quad n=1,\ldots,
$$

where  $(v_n)$  is a predictable sequence. For  $1 \leq p < \infty$  we define an operator on  $\mathcal N$  by the formula:

$$
T_p(f)=(E'|\sum v_kr_k'|^p)^{1/p},
$$

where  $f_n = \sum_{k=1}^n v_k r_k$ . Observe that  $T_p$  is a predictable operator satisfying the conditions (B1)-(B3) above, and that  $N$  is a subclass of the class of all martingales which is closed under starting and stopping (see Remark 3.2 (ii) in Hitczenko [12]). Also note that, by Fubini's theorem

$$
||T_p(f^n)||_p = ||\sum_{k=1}^n v_k r'_k||_p
$$

and by Levy's inequality we have  $T_p^*(f^n) \leq 2T_p(f^n)$ . Thus, in view of Lemma 2.2, we need to establish an appropriate upper bound on the quantity

$$
P(|f_n|\geq t||T_p(f^n)||_{\infty}),\quad t>0.
$$

In order to accomplish that goal we shall proceed in two steps. Firstly, we shall need an information on the size of the  $L_{\infty}$ -norm of  $T_p(f^n)$ . To this end it suffices to have a precise approximation for the  $L_p$ -norm of a sum  $\sum a_k r_k$ , where  $(a_k)$  is a sequence of real numbers. Unfortunately, the hitherto known two-sided estimates for  $\|\sum a_k r_k\|_p$  lack the required precision, since an upper and lower bounds usually involve constants depending on p, typically of different order of the magnitude. This difficulty will be overcome in the next section, where we shall establish a rather exact approximation for  $\|\sum a_k r_k\|_p$ . Once this is done, we shall easily obtain an appropriate exponential bound on the tail probabilities of  $\sum v_k r_k$ . This will be done in the fourth section of the present paper.

## **3. Lp-norms of Rademacher averages**

In this section we shall establish a precise approximation for the expressions of the form:

$$
\big\|\sum_{k=1}^n a_k r_k\big\|_p, \quad 1\leq p<\infty,
$$

where  $(a_k)$  is a sequence of reals. Our approach will follow a method used by Montgomery-Smith [24]. Before we state the result we need to introduce some more notation. For a sequence  $a = (a_k)$  of real numbers we let  $|a|_p =$  $(\sum |a_k|^p)^{1/p}, 1 \leq p < \infty$ . Let  $a \in \ell_2$  be a sequence of real numbers and consider the interpolation norm

$$
K(a,t;\ell_1,\ell_2)=K_{1,2}(a,t)=\inf\{|a'|_1+t|a''|_2\colon a',a''\in\ell_2,a'+a''=a\}.
$$

We shall also use the following norm on  $\ell_2$ : For  $t \in N$  and  $a \in \ell_2$  we let

$$
||a||_{P(t)} = \sup \left\{ \sum_{m=1}^{t} \left( \sum_{k \in B_m} a_k^2 \right)^{1/2} \right\},\,
$$

where the supremum is taken over all pairwise disjoint subsets,  $B_1, \ldots, B_t$  of positive integers. Those two norms are related by the following inequality established in Montgomery-Smith [24]:

LEMMA 3.1: For all  $a \in \ell_2$  and t such that  $t^2 \in N$  we have

$$
||a||_{P(t^2)} \leq K_{1,2}(a,t) \leq \sqrt{2}||a||_{P(t^2)}.
$$

We shall use this lemma to obtain the following two sided estimate on the  $L_p$ norm of Rademacher averages:

THEOREM 3.2: Let  $(r_k)$  be a Rademacher sequence. Then, there exists a con*stant c > 0 such that, for all p,*  $1 \leq p < \infty$ *, and all*  $a \in \ell_2$  *the following inequality is true:* 

$$
cK_{1,2}(a,\sqrt{p}) \leq \big\|\sum a_k r_k\big\|_p \leq K_{1,2}(a,\sqrt{p}).
$$

Proof: The right hand side inequality is a direct consequence of the Khintchine's inequality with the best constant (cf. Haagerup [9]). Given an  $\epsilon > 0$  pick a' and *a"* such that

$$
|a'|_1+\sqrt{p}|a''|_2\leq K_{1,2}(a,\sqrt{p})+\epsilon.
$$

Then we have that

$$
\|\sum a_k r_k\|_p \le \|\sum a'_k r_k\|_p + \|\sum a''_k r_k\|_p
$$
  
\n
$$
\le \sum |a'_k| + \sqrt{p} \left(\sum (a''_k)^2\right)^{1/2}
$$
  
\n
$$
\le K_{1,2}(a, \sqrt{p}) + \epsilon,
$$

and the inequality follows by letting  $\epsilon \to 0$ .

To prove the left hand side inequality we first observe that without loss of generality  $p$  can be assumed to be a positive integer. This is a consequence of the monotonicity of  $K_{1,2}(a,t)$  as a function of t and the following hypercontraction inequality for Rademacher functions due to Borell [3]:

$$
\|\sum a_k r_k\|_p \leq \sqrt{\frac{p-1}{q-1}} \|\sum a_k r_k\|_q, \quad 1 < q < p < \infty.
$$

Fix  $0 < \delta < 1$  and take  $\epsilon > 0$  such that  $(1 + \epsilon)\delta < 1$ . By Lemma 3.1 there exist pairwise disjoint subsets  $B_1, \ldots, B_p$  of N such that

$$
||a||_{P(p)} \leq (1+\epsilon) \left\{ \sum_{m=1}^p \left( \sum_{k \in B_m} a_k^2 \right)^{1/2} \right\}.
$$

$$
\|\sum a_k r_k\|_p^p = \|\sum_{m=1}^p \sum_{k \in B_m} a_k r_k\|_p^p
$$
  
\n
$$
\geq \left(\sum_{m=1}^p (1+\epsilon)\delta \left(\sum_{k \in B_m} a_k^2\right)^{1/2}\right)^p
$$
  
\n
$$
\times P\left(\bigcap_{m=1}^p \left\{\sum_{k \in B_m} a_k r_k \geq (1+\epsilon)\delta \left(\sum_{k \in B_m} a_k^2\right)^{1/2}\right\}\right)
$$
  
\n
$$
= \left(\sum_{m=1}^p (1+\epsilon)\delta \left(\sum_{k \in B_m} a_k^2\right)^{1/2}\right)^p
$$
  
\n
$$
\times \prod_{m=1}^p P\left(\sum_{k \in B_m} a_k r_k \geq (1+\epsilon)\delta \left(\sum_{k \in B_m} a_k^2\right)^{1/2}\right).
$$

By the Paley-Zygmund inequality (cf. e.g. Kahane [15, p. 24]) we infer that

$$
P\left(\sum_{k\in B_m} a_k r_k \geq (1+\epsilon)\delta\left(\sum_{k\in B_m} a_k^2\right)^{1/2}\right) \geq \frac{1}{6} \left(1-\delta^2(1+\epsilon)^2\right)^2.
$$

Therefore,  

$$
\|\sum a_k r_k\|_p^p \ge \left(\frac{1}{6} \left(1-\delta^2(1+\epsilon)^2\right)^p\right)^p \left(\sum_{m=1}^p (1+\epsilon)\delta\left(\sum_{k\in B_m} a_k^2\right)^{1/2}\right)^p,
$$

so that

$$
\|\sum a_k r_k\|_p \geq \frac{1}{6} \left(1 - \delta^2 (1 + \epsilon)^2\right)^2 (1 + \epsilon) \delta \sum_{m=1}^p \left(\sum_{k \in B_m} a_k^2\right)^{1/2}
$$

$$
\geq \frac{\delta}{6\sqrt{2}} (1 - \delta^2 (1 + \epsilon)^2)^2 K_{1,2}(a, \sqrt{p}).
$$

The left hand side inequality follows now by letting  $\epsilon \to 0$ . This completes the proof.  $\qquad \blacksquare$ 

## 4. Weak type estimate for  $T_p$

In this section we shall prove the following exponential inequality:

THEOREM 4.1: Let f be a mean-zero martingale, where  $f_n = \sum_{k=1}^n v_k r_k$ ,  $n =$ 1,..., is a *transform of a Rademacher sequence. Then,* there *exist constants*   $c_0 > 0$  and  $\kappa > 0$  such that for every  $c > c_0$  and for every positive number t the *following inequality is true:* 

$$
P(|f_n| \ge c ||K_{1,2}(v,t)||_{\infty}) \le 2 \exp \left\{-\kappa c^2 t^2\right\}.
$$

Before we pass to the proof let us observe, that Theorems 3.1 and 4.1 immediately yield the following

COROLLARY 4.2: Let  $1 \leq p < \infty$ . Under the assumptions of Theorem 4.1, there *exist constants*  $c_0 > 0$  and  $\kappa > 0$  such that for every  $c > c_0$  we have that

$$
P(|f_n| \ge c||T_p(f^n)||_{\infty}) \le 2 \exp \left\{-\kappa c^2 p\right\}, \quad n=1,\ldots.
$$

Proof of Theorem 4.1: According to Holmstedt's formula (cf. Holmstedt [14, Theorem 4.1] or Bergh and Löfström  $[2,$  Theorem 3.6.1]), for every sequence  $a \in \ell_2$  and all  $t > 0$ , the interpolation norm  $K_{1,2}(a,t)$  is equivalent (with constant not depending on  $a$  or  $t$ ) to the quantity

$$
\sum_{k=1}^{[t^2]} a^{(k)} + t \left( \sum_{k > [t^2]} (a^{(k)})^2 \right)^{1/2}
$$

Here,  $a^{(k)}$  denotes the kth largest element of the sequence  $(|a_j|)$ . We shall also write  $a_n^{(k)}$  for the kth largest element of the sequence  $(|a_1|, \ldots, |a_n|)$ , and we adopt the convention that  $a_n^{(k)} = 0$  if  $k > n$ . Let  $c_{1,2} > 0$  be any constant for which the inequality

$$
K_{1,2}(a,t) \geq c_{1,2} \left\{ \sum_{k=1}^{[t^2]} a^{(k)} + t \left( \sum_{k > [t^2]} (a^{(k)})^2 \right)^{1/2} \right\}
$$

is true for all sequences  $a \in \ell_2$  and all  $t > 0$ . Fix a positive real number t and put  $s = [t^2]$ . We shall show that the conclusion of our theorem holds with  $c_0 = 2/c_{1,2}$ and  $\kappa = c_{1,2}^2/8$ . Given a predictable sequence of random variables  $(v_n)$  let us write

$$
v_n=u_n+w_n,
$$

where

$$
u_n = \text{sgn}(v_n)(|v_n| - v_{n-1}^{(s)})^+ = \text{sgn}(v_n)(|v_n| - v_{n-1}^{(s)})I(|v_n| > v_{n-1}^{(s)}),
$$

and

$$
w_n = \mathrm{sgn}(v_n) \left( |v_n| \wedge v_{n-1}^{(s)} \right).
$$

Since  $(v_n)$  is predictable, so are  $(u_n)$  and  $(w_n)$  and we have that

$$
\sum_{k=1}^{\infty} |u_k| = \sum_{k=1}^{s} v^{(k)}
$$

and

$$
\sum_{k=1}^{\infty}w_k^2 \leq \sum_{k>s} (v^{(k)})^2
$$

(the latter follows from the fact that  $w^{(r)} \leq v^{(r+s)}$ ,  $r = 1,...$ ). Since

$$
\|\sum_{k=1}^{s} v^{(k)} + t(\sum_{k>s} (v^{(k)})^2)^{1/2}\|_{\infty} \ge \max\left\{\|\sum_{k=1}^{s} v^{(k)}\|_{\infty}, t\|(\sum_{k>s} (v^{(k)})^2)^{1/2}\|_{\infty}\right\}
$$

$$
\ge \frac{1}{2} \left\{\|\sum_{k=1}^{s} v^{(k)}\|_{\infty} + t\|(\sum_{k>s} (v^{(k)})^2)^{1/2}\|_{\infty}\right\}
$$

$$
\ge \frac{1}{2} \left\{\|\sum_{k=1}^{s} u_k\|_{\infty} + t\|(\sum_{k>s} w_k^2)^{1/2}\|_{\infty}\right\}
$$

we get that

$$
P(|\sum v_k r_k| \ge c ||K_{1,2}(v,t)||_{\infty})
$$
  
\n
$$
\le P(|\sum v_k r_k| \ge c c_{1,2} || \sum_{k=1}^{s} v^{(k)} + t (\sum_{k>s} (v^{(k)})^2)^{1/2} ||_{\infty})
$$
  
\n
$$
\le P(|\sum u_k r_k| \ge \frac{c c_{1,2}}{2} || \sum u_k ||_{\infty})
$$
  
\n
$$
+ P(|\sum w_k r_k| \ge \frac{c c_{1,2}}{2} t || (\sum w_k^2)^{1/2} ||_{\infty})
$$
  
\n
$$
\le 0 + 2 \exp \left\{-\frac{c^2 c_{1,2}^2 t^2}{8}\right\},
$$

where the last inequality follows by the subgaussian inequality for conditionally symmetric martingales (cf. e.g. Chang, Wilson and Wolff [7, Theorem 3.1] or Hitczenko [12, Lemma 4.3]), and the fact that  $cc_{1,2}/2 > 1$ . This completes the proof.  $\blacksquare$ 

## **5. Domination inequality**

The inequalities proved in the preceding two sections immediately imply the following result:

THEOREM 5.1: There exists an absolute constant K such that, for every p,  $1 \leq$  $p < \infty$ , and every martingale f with difference sequence of the form:  $d_n = v_n r_n$ ,  $n = 1, \ldots$  we have

$$
||f_n^*||_p \le K \big\| \sum_{k=1}^n v_k r'_k \big\|_p, \quad n = 1, \ldots
$$

*Remark:* The above theorem remains true if Rademacher sequence is replaced by a sequence  $(\gamma_k)$  of i.i.d. standard Gaussian variables. This follows from the fact that

$$
(E'|\sum v_k\gamma'_k|^p)^{1/p} = \|\gamma_1\|_p (\sum v_k^2)^{1/2} \approx \sqrt{p}(\sum v_k^2)^{1/2}
$$

and from the estimate

$$
P(|\sum v_k \gamma_k| \geq t) \leq 2 \exp\left\{\frac{-t^2}{2\|\sum v_k^2\|_{\infty}}\right\}.
$$

Theorem 5.1 can be easily extended to martingale transforms of sequences satisfying some regularity conditions (although with constant  $K$  depending on "regularity constant" of that sequence). Following Gundy [8], we say that a sequence  $({\xi_k})$  is  $L^{\infty}$ -regular if  ${\xi_k}$ 's are uniformly bounded and there exists a  $\delta > 0$  such that, for each  $k = 1, \ldots,$ 

$$
E|\xi_k| \geq \delta \|\xi_k\|_{\infty}.
$$

With the above definition we have:

THEOREM 5.2: Let  $(\xi_n)$  be  $L^\infty$ -regular sequence of independent mean-zero ran*dom variables. Then, there exists a constant K depending only on*  $\delta$  *such that,* for every predictable sequence  $(v_k)$  we have that

$$
\big\|\sum v_k \xi_k\big\|_p\leq K\big\|\sum v_k \xi_k'\big\|_p,\quad 1\leq p<\infty.
$$

Proof: Let us observe first, that if  $\xi$  and  $\xi'$  are i.i.d. mean-zero random variables such that  $||\xi||_1 \geq \delta ||\xi||_{\infty}$  then

$$
\|\xi-\xi'\|_1\geq \|\xi\|_1\geq \delta \|\xi\|_{\infty}\geq \frac{\delta}{2}\|\xi-\xi'\|_{\infty}.
$$

Moreover

$$
\big\| \sum v_k \xi_k \big\|_p \le \big\| \sum v_k \xi_k' \big\|_p + \big\| \sum v_k (\xi_k - \xi_k') \big\|_p.
$$

Therefore, we can and do assume in the proof that  $\xi_k$ 's are symmetric. We shall now use an observation due to Wang [25, the bottom half of p. 399] (cf. also Hitczenko [12, proof of Lemma 4.3]). Define a sequence of  $\sigma$ -algebras  $(\mathcal{A}_n)$  by letting  $A_n = \sigma(\mathcal{F}_n, |\xi_{n+1}|)$ . It follows from the symmetry that, relative to  $(A_n)$ ,  $(v_n\xi_n)$  is a martingale difference sequence equidistributed with  $(u_ne_n)$ , where  $(\epsilon_n)$  is a Rademacher sequence and  $(u_n) = (v_n|\xi_n|)$  is predictable sequence. Let  $(\epsilon'_n)$  be a Rademacher sequence independent of all other random variables under consideration. We shall write

$$
T_p^{\xi}(f) = \left(E' \big| \sum v_k \xi'_k \big|^p \right)^{1/p}
$$

and

$$
T_p^{\epsilon}(f) = \left(E' \big| \sum v_k \epsilon'_k \big|^p \right)^{1/p}.
$$

With this notation, denoting by  $E'_{\epsilon}$  and  $E'_{\epsilon}$  integration with respect to  $(\epsilon'_{n})$  and  $({\xi}_{n}^{\prime})$ , respectively we have

$$
T_p^{\xi}(f) = \left(E_{\xi}' | \sum v_k \xi_k' |^p\right)^{1/p} = \left(E_{\epsilon}' E_{\xi}' | \sum v_k | \xi_k' | \epsilon_k' |^p\right)^{1/p}
$$
  
\n
$$
\geq \left(E_{\epsilon}' | \sum v_k ||\xi_k||_1 \epsilon_k' |^p\right)^{1/p} \geq \delta \left(E_{\epsilon}' | \sum v_k ||\xi_k||_{\infty} \epsilon_k' |^p\right)^{1/p}
$$
  
\n
$$
\geq \delta \left(E_{\epsilon}' | \sum v_k |\xi_k| \epsilon_k' |^p\right)^{1/p} = \delta T_p^{\epsilon}(f),
$$

where the last inequality follows by the contraction principle (cf. e.g. Ledoux and Talagrand [22, Theorem 4.4]). Therefore,

$$
||T_p^{\xi}(f^n)||_{\infty}\geq \delta||T_p^{\epsilon}(f^n)||_{\infty},
$$

so that, for each  $c > 0$  we have

$$
P(f^* \ge c||T_p^{\xi}(f)||_{\infty}) \le P(f^* \ge c\delta||T_p^{\epsilon}(f)||_{\infty}).
$$

The result now follows from Lemma 2.2 and Corollary 4.2.  $\blacksquare$ 

## 6. An **example**

It is natural to ask whether the results of the preceding section (Theorem 5.1 in particular) can be strengthened to domination of tail probabilities. The answer to this question turns out to be negative. We would like to conclude by presenting an example, which was communicated to us by M. Talagrand and which is included here with his kind permission.

THEOREM 6.1: There *is no constant C which* makes *the inequality:* 

$$
P(|\sum v_j r_j| \ge Ct) \le CP(|\sum v_j r_j' | \ge t)
$$

*true.* 

Proof: Given an integer  $k$  let  $N_1$  be an integer which will be specified later and let  $N_2, \ldots N_k$  be defined by:

$$
N_i - N_{i-1} = 2^{-(i-1)} N_1, \quad i = 2, \ldots, k.
$$

Put

$$
\Omega_1 = \{r_1 = \cdots = r_{N_1} = 1\},\
$$

and then

$$
\Omega_i = \Omega_{i-1} \cap \{r_{N_{i-1}+1} = \cdots = r_{N_i}\} \quad i=2,\ldots,k.
$$

Define a predictable sequence of random variables  $(v_i)$  by the formulas:

$$
v_1 = \dots = v_{N_1} = 1
$$
  
\n
$$
v_{N_1+1} = \dots = v_{N_2} = 2I_{\Omega_1}
$$
  
\n
$$
\dots \dots \dots
$$
  
\n
$$
v_{N_{k-1}+1} = \dots = v_{N_k} = 2^{k-1}I_{\Omega_{k-1}}.
$$

Then for  $t = N_1$  we have:

$$
P(|\sum_{j=1}^{N_k} v_j r_j| \ge kt) = 2P(r_j = 1, j = 1, ..., N_k) = 2 \cdot 2^{-N_1(1+2^{-1}+...+2^{-(k-1)})}.
$$

We shall now estimate

$$
P\Big(\Big|\sum_{j=1}^{N_{\boldsymbol{k}}}v_jr'_j\Big|\geq 4t\Big)
$$

**from above.** 

$$
P\left(\left|\sum_{j=1}^{N_{k}} v_{j} r_{j}'\right| \geq 4t\right) = P\left(\left|\sum_{j=1}^{N_{k}} v_{j} r_{j}'\right| \geq 4t, \Omega_{1}^{c}\right)
$$
  
+
$$
P\left(\left|\sum_{j=1}^{N_{k}} v_{j} r_{j}'\right| \geq 4t, \Omega_{1} \setminus \Omega_{2}\right) + \cdots
$$
  
+
$$
P\left(\left|\sum_{j=1}^{N_{k}} v_{j} r_{j}'\right| \geq 4t, \Omega_{k-1} \setminus \Omega_{k}\right)
$$
  
+
$$
P\left(\left|\sum_{j=1}^{N_{k}} v_{j} r_{j}'\right| \geq 4t, \Omega_{k}\right)
$$
  
=
$$
P\left(\left|\sum_{j=1}^{N_{1}} r_{j}'\right| \geq 4t\right) (1 - P(\Omega_{1}))
$$
  
+
$$
P\left(\left|\sum_{j=1}^{N_{1}} r_{j}' + 2 \cdot \sum_{j=N_{1}+1}^{N_{2}} r_{j}'\right| \geq 4t\right) P(\Omega_{1} \setminus \Omega_{2}) + \cdots
$$
  
+
$$
P\left(\left|\sum_{j=1}^{N_{1}} r_{j}' + \cdots + 2^{k-1} \cdot \sum_{j=N_{k-1}+1}^{N_{k}} r_{j}'\right| \geq 4t\right) P(\Omega_{k-1}).
$$

Since  $4t = 4N_1 > N_1 + 2(N_2 - N_1) + 4(N_3 - N_2)$ , the first three terms in the last sum are zero, and the ith term can be estimated as follows:

$$
P(\Omega_{i-1} \setminus \Omega_i) \le P(\Omega_{i-1}) = 2^{-N_{i-1}} = 2^{-N_1(1+\cdots+2^{-(i-2)})}.
$$

By subgaussian inequality for Rademacher functions, one has:

$$
P\Big(\Big|\sum_{j=1}^{N_1} r'_j + \dots + 2^{i-1} \cdot \sum_{j=N_{i-1}+1}^{N_i} r'_j\Big| \ge 4t\Big)
$$
  

$$
\le 2 \exp\left\{\frac{-16t^2}{2E(\sum_{j=1}^{N_1} r'_j + \dots + 2^{i-1} \sum_{j=N_{i-1}+1}^{N_i} r'_j)^2}\right\}.
$$

**But** 

$$
E\Big(\sum_{j=1}^{N_1} r'_j + \cdots + 2^{i-1} \cdot \sum_{j=N_{i-1}+1}^{N_i} r'_j\Big)^2 = N_1(1+\cdots+2^{i-1}) \leq 2^i N_1,
$$

so that the exponent above is less or equal than

$$
\exp\{-8N_1/2^i\} \leq 2^{-8N_1/2^i}.
$$

Combining the above estimates we obtain that:

$$
P\Big(\Big|\sum_{j=1}^{N_1} r'_j + \dots + 2^{i-1} \cdot \sum_{j=N_{i-1}+1}^{N_i} r'_j\Big| \ge 4t\Big) P(\Omega_{i-1})
$$
  
\n
$$
\le 2 \cdot 2^{-8N_1/2^i} \cdot 2^{-N_1(1+\dots+2^{-(i-2)})}
$$
  
\n
$$
= 2 \cdot 2^{-N_1(1+\dots+2^{-(k-1)})} \cdot 2^{-8N_1/2^i} \cdot 2^{N_1(2^{-(i-1)}+\dots+2^{-(k-1)})}
$$
  
\n
$$
\le P\Big(\Big|\sum_{j=1}^{N_k} v_j r_j\Big| \ge kt\Big) \cdot 2^{-8N_1/2^i} \cdot 2^{N_1/2^{i-2}}
$$
  
\n
$$
= 2^{-N_1/2^{i-2}} P\Big(\Big|\sum_{j=1}^{N_k} v_j r_j\Big| \ge kt\Big).
$$

Therefore, summing over  $i$  we get that

$$
P\Big(\Big|\sum_{j=1}^{N_k} v_j r'_j\Big| \ge 4t\Big) \le \sum_{i=4}^k 2^{-N_1/2^{i-2}} P\Big(\Big|\sum_{j=1}^{N_k} v_j r_j\Big| \ge kt\Big) \le k \cdot 2^{-N_1/2^{k-2}} P\Big(\Big|\sum_{j=1}^{N_k} v_j r_j\Big| \ge kt\Big).
$$

This completes the proof since  $N_1$  can be made arbitrarily large.

Remark: It is probably worth noticing that the above example shows a little bit more, namely that  $v_k$ 's may be chosen to be of the form  $v_k = a_k I(\tau \ge k)$ ,  $k = 1, \ldots$ , where  $(a_k)$  is a nonrandom sequence and  $\tau$  is a stopping time; one just puts

$$
a_j = 2^{i-1}
$$
 for  $N_{i-1} < j \le N_i$ ,  $i = 1, ..., k$ ;

and  $\tau = \inf\{n : r_n \neq r_{n-1}\}.$  This shows that moment inequality proved by Klass [16, Theorem 3.1] for randomly stopped sums of independent random variables can not be strengthened to distributional statement, even for linear combinations of Rademacher functions, l

ACKNOWLEDGEMENT: It is a pleasure to thank Michel Talagrand who generously permitted me to include his example (Theorem 6.1) in this paper. I would also like to thank the referee for pointing out several omissions.

#### **References**

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